

Correcting biased observation model error in data assimilation

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Joint work with John Harlim, PSU

BIAS IN OBSERVATION MODELS

- ▶ Consider the standard filtering problem,

$$\begin{aligned}x_i &= f(x_{i-1}) + \omega_{i-1} \\y_i &= h(x_i) + \eta_i\end{aligned}$$

- ▶ We assume the true observation function $h(x)$ is unknown
- ▶ An approximate model is available $\tilde{h}(x)$ so that

$$y_i = h(x_i) + \eta_i = \tilde{h}(x_i) + b_i + \eta_i$$

- ▶ Where $b_i \equiv h(x_i) - \tilde{h}(x_i)$ is called the bias

EXAMPLE 1: LORENZ-96

- ▶ Consider the standard 40-dimensional Lorenz-96,

$$\dot{x}_j = x_{j-1}(x_{j+1} - x_{j-2}) - x_j + 8$$

- ▶ We observe 20 of the 40 variables
- ▶ We draw $\xi_i \sim \mathcal{U}(0, 1)$ and let the observations be,

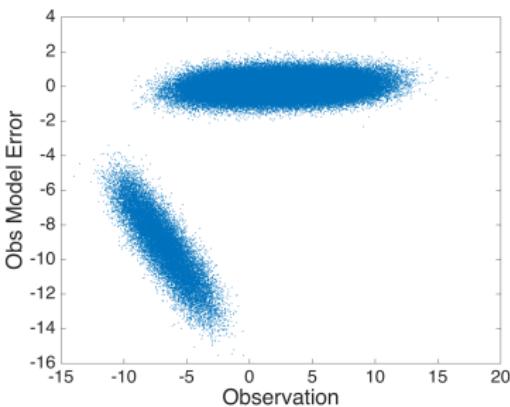
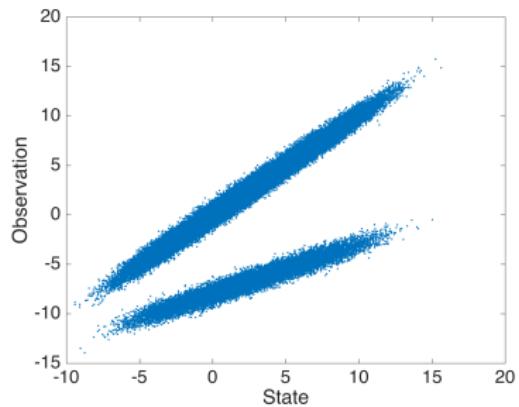
$$h(x_k) = \begin{cases} x_k & \xi_i > 0.8 \\ \beta_k x_k - 8 & \text{else} \end{cases}$$

$$\beta_k \sim \mathcal{N}(0.5, 1/50).$$

- ▶ h is applied to 7 randomly chosen variables
- ▶ Remaining 13 are directly observed

EXAMPLE 1: LORENZ-96

- ▶ The result is a bimodal distribution, “cloudy/clear”
- ▶ Obs Model Error = True Obs - $\tilde{h}(\text{True State})$



CORRECTING THE BIAS

- ▶ Our goal is to find $p(b_i | y_i)$
- ▶ We can then adjust the filter by defining a new innovation

$$\hat{\epsilon}_i = \epsilon_i + \hat{b}_i = y_i - \tilde{h}(x_i^f) + \hat{b}_i$$

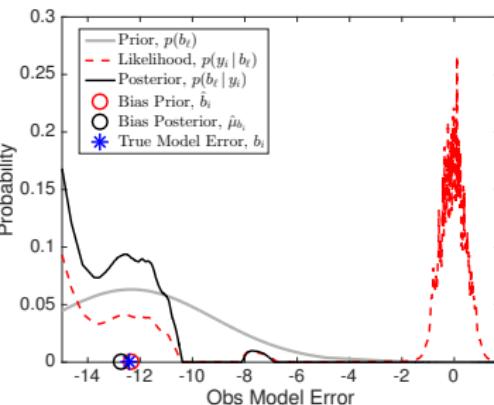
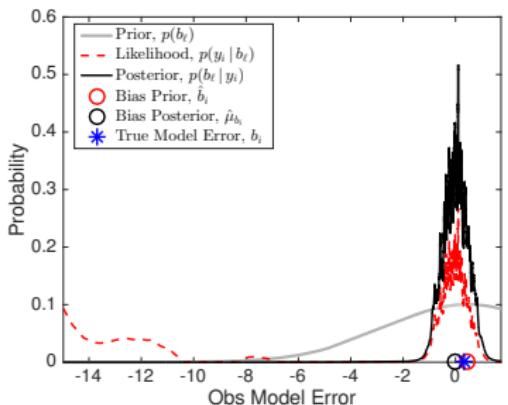
- ▶ Where $\hat{b}_i = \mathbb{E}_{p(b_i | y_i)}[b_i]$
- ▶ We also inflate the obs covariance by $R_i = R^o + \hat{R}_{b_i}$
- ▶ Where $\hat{R}_{b_i} = \mathbb{E}_{p(b_i | y_i)}[(b_i - \hat{b}_i)(b_i - \hat{b}_i)^\top]$

CORRECTING THE BIAS

- ▶ If we can estimate $p(b_i | y_i)$ we can ‘fix’ the obs
- ▶ We will use Bayes’ to find $p(b_i | y_i) = p(b_i)p(y_i | b_i)$
- ▶ We will use a simple prior $p(b_i) = \mathcal{N}(\epsilon_i, P_i^y)$
 - ▶ $\epsilon_i = y_i - h(x_i^f)$ is the innovation
 - ▶ P_i^y is the innovation covariance estimate
- ▶ The real challenge is to estimate $p(y_i | b_i)$
- ▶ We will learn $p(y_i | b_i)$ from training data using the kernel estimation of conditional distributions

CORRECTING THE BIAS

- ▶ Below plots have $y_i \approx -4$
- ▶ Left is clear, right is cloudy
- ▶ Notice bimodal likelihood



LEARNING THE CONDITIONAL DISTRIBUTION

- Given training data (y_i, b_i) our goal is to learn $p(y_i | b_i)$
- For a kernel $K(\alpha, \beta) = e^{-\frac{||\alpha - \beta||^2}{\delta^2}}$ we define Hilbert spaces

$$\mathcal{H}_y = \left\{ \sum_{i=1}^N a_i K(y_i, \cdot) : \vec{a} \in \mathbb{R}^N \right\}, \quad \mathcal{H}_b = \left\{ \sum_{i=1}^N a_i K(b_i, \cdot) : \vec{a} \in \mathbb{R}^N \right\}$$

- For example the kernel density estimate (KDE) \hat{q} is in \mathcal{H}_y

$$\hat{q}(y) = \frac{1}{m_0 N} \sum_{i=1}^N K(y_i, y)$$

- Eigenvectors ϕ_ℓ of $K_{ij} = K(y_i, y_j)$ form an orthonormal basis for \mathcal{H}_y . Similarly φ_k are a basis for \mathcal{H}_b .

LEARNING THE CONDITIONAL DISTRIBUTION

- ▶ We assume that $p(y | b)$ can be approximated in $\mathcal{H}_y \otimes \mathcal{H}_b$
- ▶ Let $C_{ij}^{yb} = \langle \phi_i, \varphi_j \rangle$ and $C_{ij}^{bb} = \langle \varphi_i, \varphi_j \rangle$ then define

$$C^{y|b} = C^{yb} \left(C^{bb} + \lambda I \right)^{-1}$$

- ▶ We can then define a consistent estimator of $p(y | b)$ by

$$\hat{p}(y | b) = \sum_{i,j=1}^N C_{i,j}^{y|b} \phi_i(y) \varphi_j(b) \hat{q}(y)$$

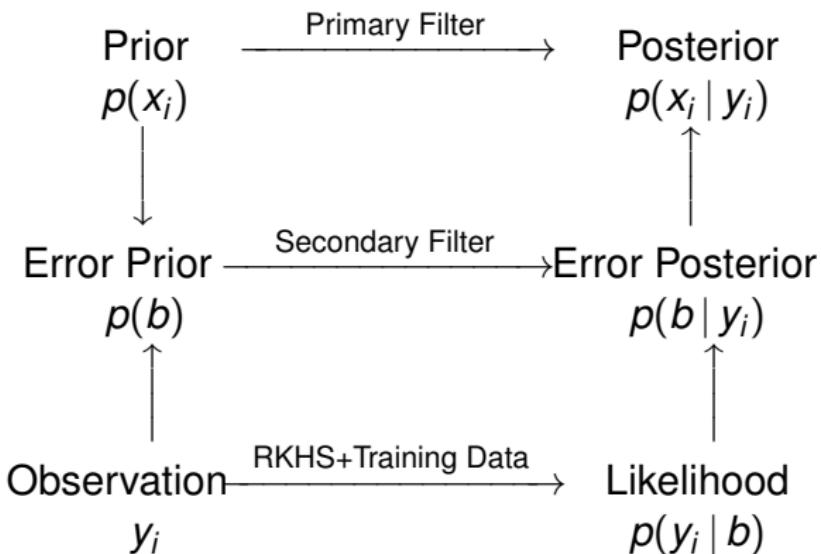
- ▶ We define eigenfunctions with Nystöm extension

$$\varphi_j(b) = \lambda_j^{-1} \sum_{i=1}^N \varphi_j(b_i) K(b_i, b)$$

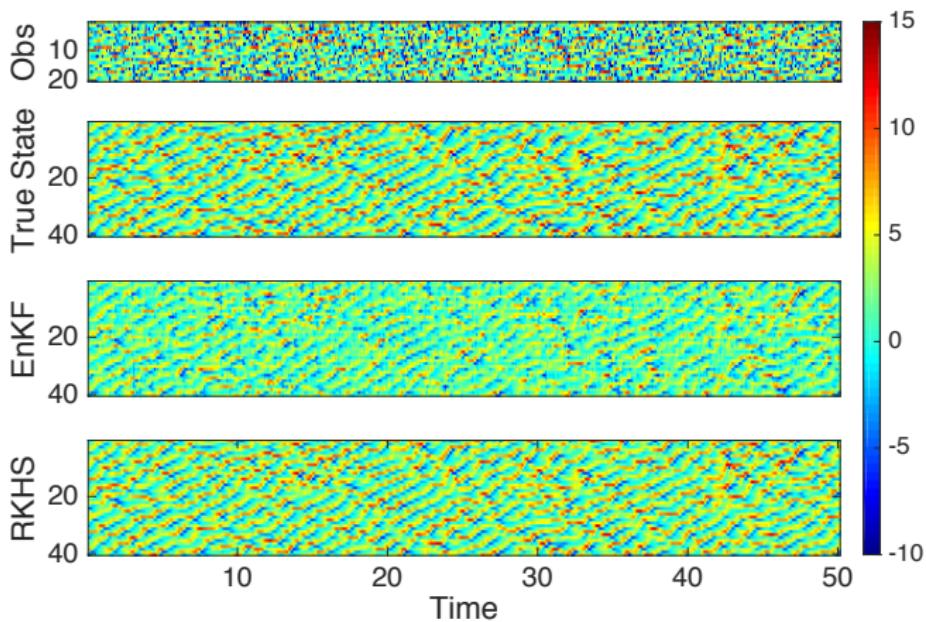
OVERVIEW

- ▶ **Learning Phase:** Given training data set (x_i, y_i)
 - ▶ Compute the biases $b_i = y_i - \tilde{h}(x_i)$
 - ▶ Learn the conditional distribution $p(y | b)$
- ▶ **Filtering:** Forecast $x_i^f \Rightarrow$ innovation $\epsilon_i = y_i - \tilde{h}(x_i^f)$
- ▶ Use prior $p(b) = \mathcal{N}(\epsilon_i, P_i^y)$
- ▶ Combine with conditional to find $p(b | y_i) = p(b)p(y_i | b)$
- ▶ Estimate conditional mean \hat{b}_i and covariance \hat{R}_{b_i}
- ▶ Adjust innovation $\hat{\epsilon}_i = \epsilon_i + \hat{b}_i$ and $R_i = R^o + \hat{R}_{b_i}$
- ▶ Apply Kalman update, continue to the next filter step

OVERVIEW

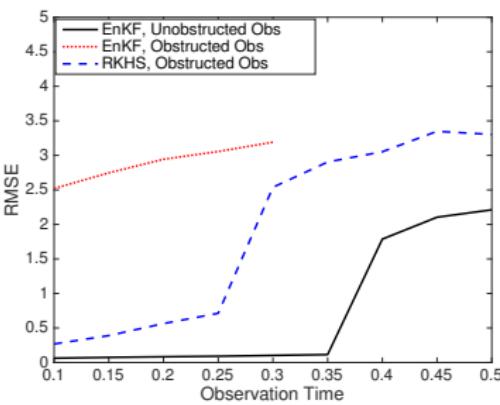
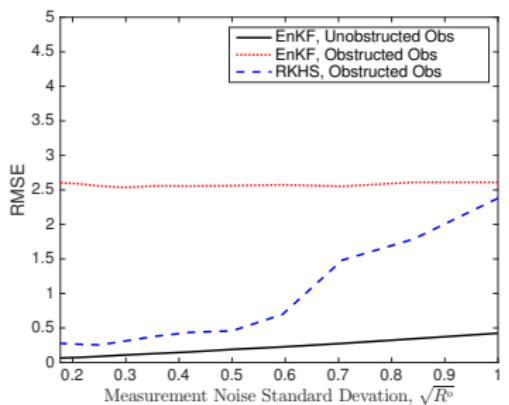


LORENZ-96 RESULTS



LORENZ-96 RESULTS

- ▶ Works well with small measurement noise
- ▶ Observations need to be precise, but not accurate



EXAMPLE 2: MULTI-CLOUD “SATELLITE-LIKE” OBS

- ▶ Consider a 7-dim'l model for a column of atmosphere
 - ▶ Baroclinic anomaly potential temperatures, θ_1 and θ_2
 - ▶ Boundary layer anomaly potential temperature, θ_{eb}
 - ▶ Vertically averaged water vapor content, q
 - ▶ Cloud fractions: congestus f_c , deep f_d , and stratiform f_s
- ▶ Extrapolate anomaly potential temperature at height z

$$T(z) = \theta_1 \sin\left(\frac{z\pi}{Z_T}\right) + 2\theta_2 \sin\left(\frac{2z\pi}{Z_T}\right), \quad z \in [0, 16]$$

Khouider, B., J. Biello, and A. J. Majda, 2010: A stochastic multicloud model for tropical convection.

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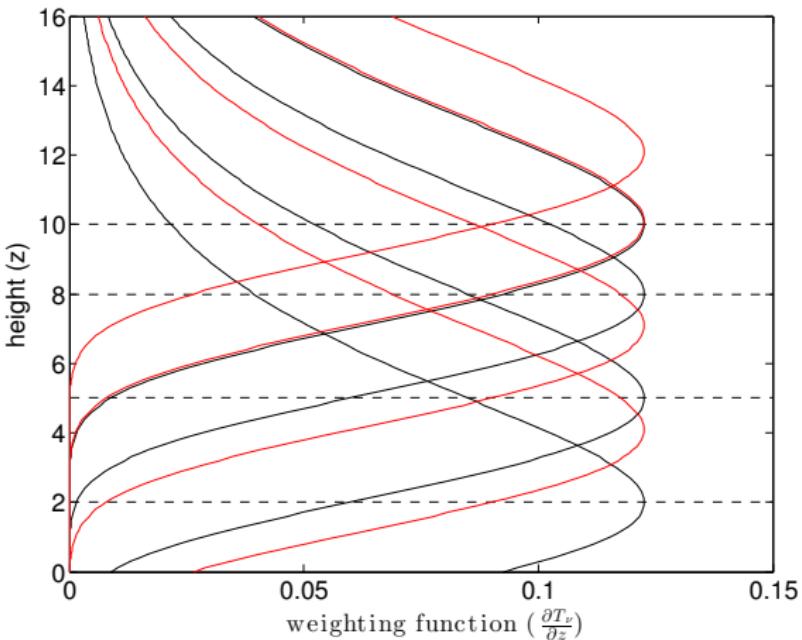
- ▶ Brightness temperature-like quantity at wavenumber- ν

$$\begin{aligned} h_\nu(x, f) = & (1 - f_d - f_s) \left[(1 - f_c)(\theta_{eb} T_\nu(0) + \int_0^{z_c} T(z) \frac{\partial T_\nu}{\partial z}(z) dz) \right. \\ & + f_c T(z_c) T_\nu(z_c) + \int_{z_c}^{z_d} T(z) \frac{\partial T_\nu}{\partial z}(z) dz \Big] \quad (1) \\ & + (f_d + f_s) T(z_d) T_\nu(z_d) + \int_{z_d}^{\infty} T(z) \frac{\partial T_\nu}{\partial z}(z) dz, \end{aligned}$$

- ▶ Setting $f = 0$ is the clear sky model

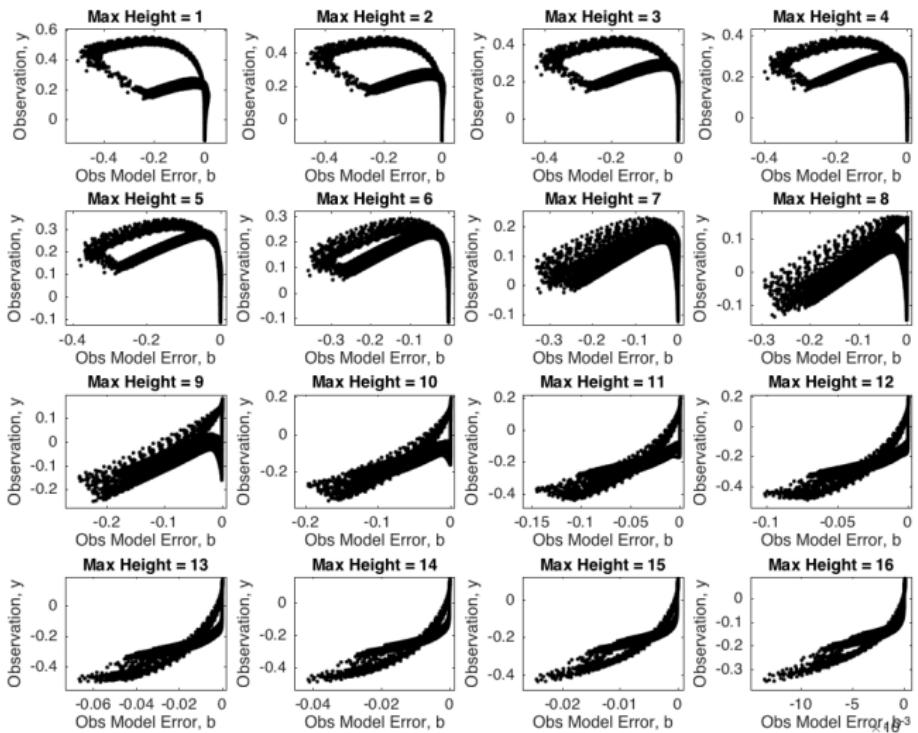
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- Weighting functions define RTM at different wavenumbers



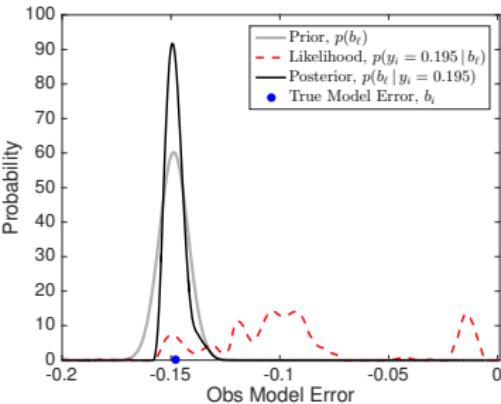
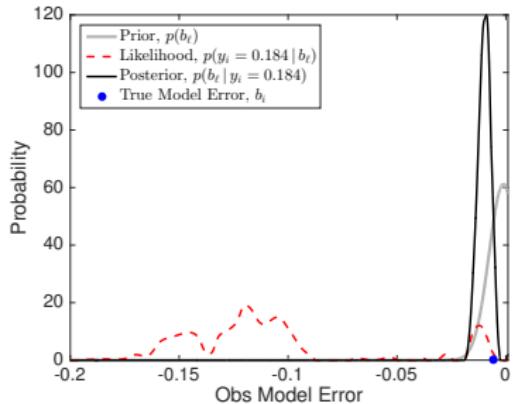
EXAMPLE 2: MULTI-CLOUD “SATELLITE-LIKE” OBS

- ▶ Biases at the 16 observed wavenumbers

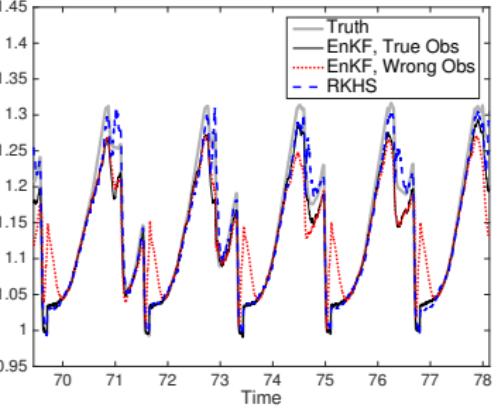
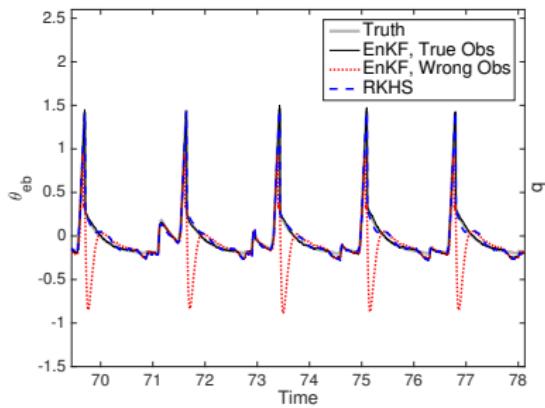
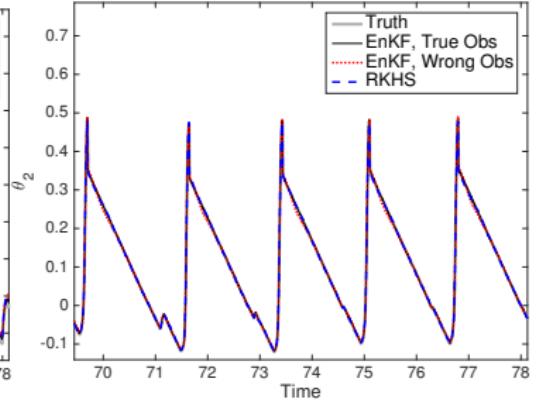
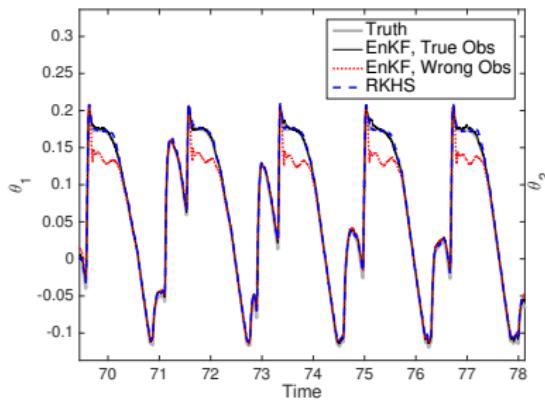


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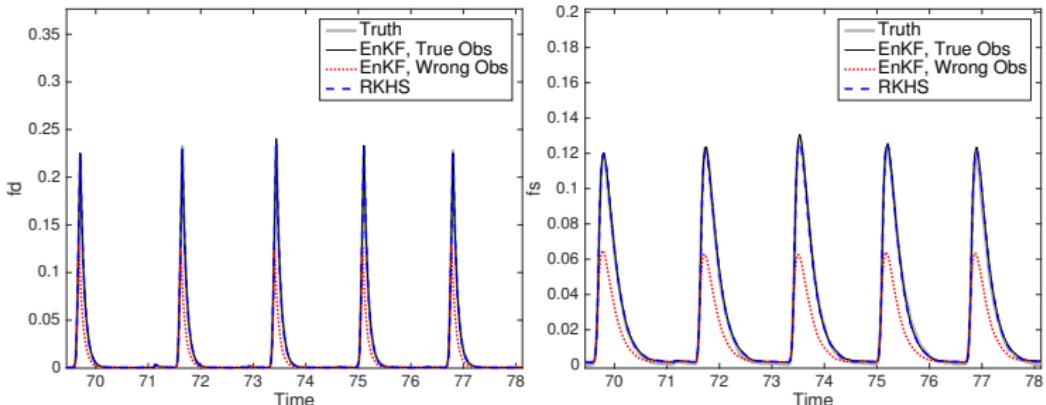
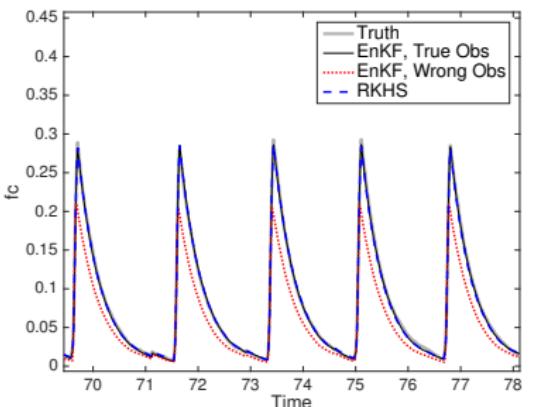
► Multimodal likelihood functions



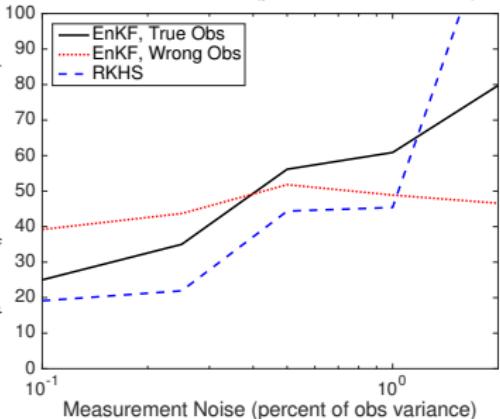
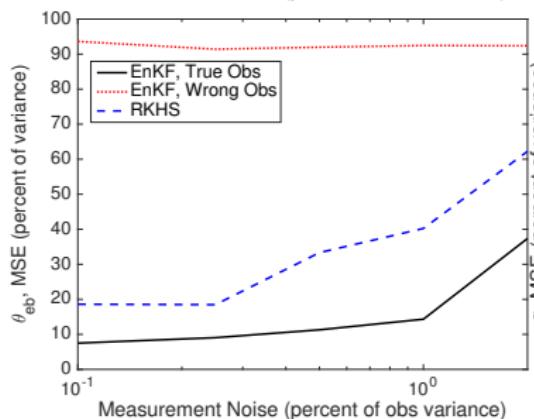
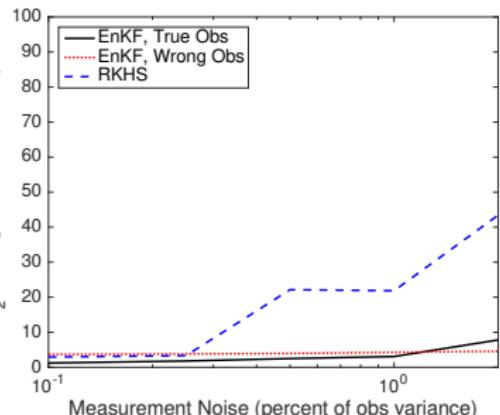
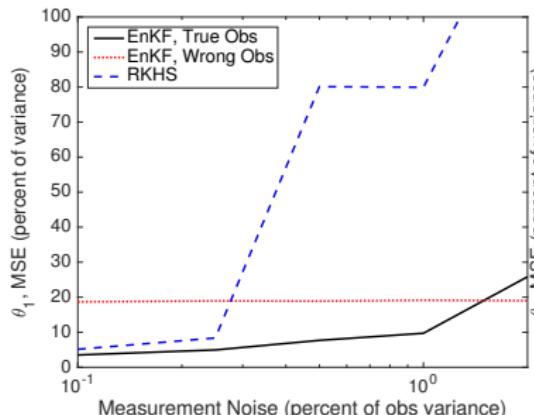
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